

1 **A framework for assessing the skill and value of operational recruitment**

2 **forecasts**

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## 10 **Abstract**

11 Forecasting variation in the recruitment to fish stocks is one of the most challenging and long-running  
12 problems in fisheries science and essentially remains unsolved today. Traditionally recruitment forecasts  
13 are developed and evaluated based on explanatory and goodness-of-fit approaches that do not reflect  
14 their ability to predict beyond the data on which they were developed. Here we propose a new generic  
15 framework that allows the skill and value of recruitment forecasts to be assessed in a manner that is  
16 relevant to their potential use in an operational setting. We assess forecast skill based on predictive power  
17 using a retrospective forecasting approach inspired by meteorology, and emphasise the importance of  
18 assessing these forecasts relative to a baseline. We quantify the value of these forecasts using an  
19 economic cost-loss decision model that is directly relevant to many forecast users. We demonstrate this  
20 framework using four stocks of lesser sandeel (*Ammodytes marinus*) in the North Sea, showing for the  
21 first time in an operationally realistic setting that skilful and valuable forecasts are feasible in two of these  
22 areas. This result shows the ability to produce valuable short-term recruitment forecasts, and highlights  
23 the need to revisit our approach to and understanding of recruitment forecasting.

## 24 Introduction

25 Recent developments in ocean observations and modelling today make it possible to forecast many of the  
26 physical variables in the ocean (Doblas-Reyes *et al.*, 2013; Meehl *et al.*, 2014). Building on top of this data  
27 about the ocean environment, forecasts of marine ecological responses have been developed (Payne *et*  
28 *al* 2017) and provide managers and stakeholders the foresight needed to sustainably manage marine living  
29 resources (Tommasi *et al.*, 2017b; Hobday *et al.*, 2018). Examples of operational forecasts already in use  
30 include southern Bluefin tuna habitat forecasts (Eveson *et al.*, 2015), the dynamic fisheries bycatch  
31 management tool EcoCast (Hazen *et al.*, 2018), and blue whale habitat preference forecast (Hazen *et al.*,  
32 2017). However, these operational fisheries forecast products are currently limited to predictions of  
33 distribution and phenology and there are currently no known operational marine fish recruitment  
34 forecasts (Payne *et al.*, 2017).

35 Understanding and forecasting changes in fish stock productivity has, however, been a key aspiration in  
36 fisheries science for the last century (Leggett and DeBlois, 1994; Subbey *et al.*, 2014; Tommasi *et al.*,  
37 2017a; Haltuch *et al.*, 2019). Recruitment, the number of young individuals produced each year, has a key  
38 role in shaping fish population dynamics (Hilborn and Walters, 1992), especially in determining total  
39 allowable catches for short-lived species, where the recruiting year-classes contribute a significant share  
40 of the landings. Environmental drivers play an important role in shaping the productivity of such stocks  
41 (e.g. via temperature (MacKenzie *et al.*, 2008; Mantzouni and Mackenzie, 2010), salinity (Köster *et al.*,  
42 2005) or phenology (Platt *et al.*, 2003)) and including climate information in stock-assessments can reduce  
43 uncertainties in stock status and the risk of over- or under harvesting (Hare *et al.*, 2010; Haltuch and Punt,  
44 2011; Tommasi *et al.*, 2017a, 2017b). The ability to foresee changes in productivity on a short time-scale  
45 can therefore enable adaptive and pre-emptive decision-making strategies, benefiting both stakeholders  
46 and managers (Hobday *et al.*, 2016; Payne *et al.*, 2017; Welch *et al.*, 2019).

47 Common approaches have however shown limited ability to produce reliable recruitment forecasts for  
48 operational (i.e. regularly repeated) use in management. The large variety of underlying environmental,  
49 physical and ecosystem processes affecting recruitment simultaneously (Leggett and DeBlois, 1994;  
50 Browman *et al.*, 1995; Myers, 1998; Tommasi *et al.*, 2017b) can often give rise to transient but spurious  
51 correlations (Sugihara *et al.*, 2012). Fish population time series are often relatively short in length (Ricard  
52 *et al.*, 2012) and hampered by high observation noise, limiting the ability to develop and test predictive  
53 models (Clark and Bjørnstad, 2004; Ward *et al.*, 2014). Furthermore, environment-recruitment  
54 correlations have been shown to breakdown when confronted with new data, diminishing the uses for  
55 management (Myers, 1998; Tommasi *et al.*, 2017b). The relative importance of drivers of recruitment  
56 can also change from year to year (“non-stationarity”) (Subbey *et al.*, 2014; Haltuch *et al.*, 2019). As a  
57 consequence of all of these processes, recruitment forecasts are widely viewed with scepticism in the  
58 community today.

59 Nevertheless, the potential of such forecasts to benefit all those that depend on living marine resources  
60 is clear. So how can this potential be realised? And even more importantly, how would we know when we  
61 have produced forecasts that can be used as a regular part of decision-making? To answer this question,  
62 here we take inspiration from other forecasting fields, and in particular from meteorology, a discipline  
63 that has also been attempting to predict chaotic and difficult to observe systems for nearly a century  
64 (albeit with considerably more success!). In particular, the question of “what makes a good forecast?” is  
65 addressed in a seminal 1993 paper in the field by Alan Murphy (Murphy, 1993) that introduces two key  
66 relevant concepts, skill, and value, which form the basis for this work.

67 Murphy defines forecast “skill” as the quantitative ability of the forecast: is it numerically correct? In the  
68 marine setting, model performance is often measured based on goodness-of-fit measures that quantify  
69 the ability to explain the data e.g (Lindegren *et al.*, 2018). There is however, a fundamental difference  
70 between explanatory and predictive power: while *explanatory* models can be used to investigate causal

71 hypotheses, models with high explanatory power cannot be expected to predict well (Levins, 1966;  
72 Shmueli, 2009; Dickey-Collas et al., 2014). But when the goal is to produce forecasts to be used regularly  
73 to predict into the future for use in a decision-making context, we clearly need to evaluate their *predictive*  
74 power. In the atmospheric and climate sciences for example, skill is often assessed based on retrospective  
75 forecast analysis (Wilks, 2011) i.e. predicting beyond the period over which the model was developed or  
76 tuned, directly reflecting the way the forecast would be used operationally. Furthermore, meteorology  
77 always places its forecasts in the context of a baseline or reference forecast (Jolliffe and Stephenson, 2012;  
78 Payne *et al.*, 2012). Common baseline forecasts includes random selection of categories or using the  
79 average over a given reference time period, often referred to as climatology in atmospheric sciences  
80 (Jolliffe and Stephenson, 2012).

81 Secondly, Murphy discusses the usefulness of a forecast in terms of its “value” in aiding decision-making.  
82 A good forecast is of value to an end-user by assisting in decision-making, providing economic value or  
83 otherwise benefiting the user (Murphy, 1993). While value in recruitment forecasts has been discussed  
84 (e.g. Walters, 1989; Field *et al.*, 2010), a quantitative approach to value is rarely seen in marine science.  
85 Simple economic decision models can analyse forecasts under simplified assumptions, helping end-users  
86 decide if it is economically wise to follow the forecast (Murphy, 1976a). Quantitatively providing a value  
87 assessment can help integrate forecast products directly into a user’s framework, allowing users to assess  
88 the benefits of a given forecast system and can give a clear insight into how, and when, a forecast should  
89 be used (Murphy, 1976a).

90 Here we argue that as the recruitment problem has never been evaluated from this perspective before,  
91 we currently do not know whether it is possible to regularly make skilful and valuable forecasts of  
92 recruitment. We therefore combine the ideas Murphy (1993) with the state of the art in recruitment  
93 modelling to give a generic framework for developing and assessing short-term recruitment forecasts for  
94 fish stocks for regular use in an decision-making setting. Forecast skill is assessed based on predictive

95 performance, using validation techniques currently used in atmospheric and meteorological sciences and  
96 that reflect the way a forecast would be used in practice. Value is assessed quantitatively, using an economic  
97 cost-loss decision model, providing insight into the actual monetary value of the forecast product. We  
98 demonstrate the framework using multiple stocks of the ecologically and economically important lesser  
99 sandeel (*Ammodytes marinus*) in the North Sea, where previous studies of recruitment have already  
100 highlighted several recruitment correlates (Arnott and Ruxton, 2002; van Deurs *et al.*, 2009; Lindegren *et*  
101 *al.*, 2018).

## 102 **Methods**

### 103 **Recruitment forecast framework**

104 This work presents a generic framework (*Figure 1a*) for assessing recruitment forecasts of fish stocks in an  
105 operational setting. The core of the framework is the idea of retrospective forecasting, an approach  
106 adapted from the atmospheric sciences, in which the time series of interest is split into two continuous  
107 blocks either side of a hypothetical “forecast issue date”. The first block is used to parameterise and train  
108 the core predictive model (the “training” block): predictions are then made for the remaining block of  
109 data (the “verification” block) based on this model. The issue date is then shifted forward by one time  
110 step, the data repartitioned and the process repeated. Iterating over all issue dates, a database of  
111 predictions is generated, with each prediction being characterised by the id of the cohort being predicted  
112 and issue date: the difference between these two is the “lead time” of the forecast (*Figure 1b*). The  
113 ensemble of predictions can then be compared against the “true” recruitment to that cohort, with various  
114 skill metrics being calculated as a function of forecast lead time. The skill metrics generated are then used  
115 as the basis for forecast value assessment.

116 There are several key features of this framework that make it highly appropriate for addressing the  
117 question at hand i.e. assessing operational forecast skill. The emphasis on temporal blocks, for example,

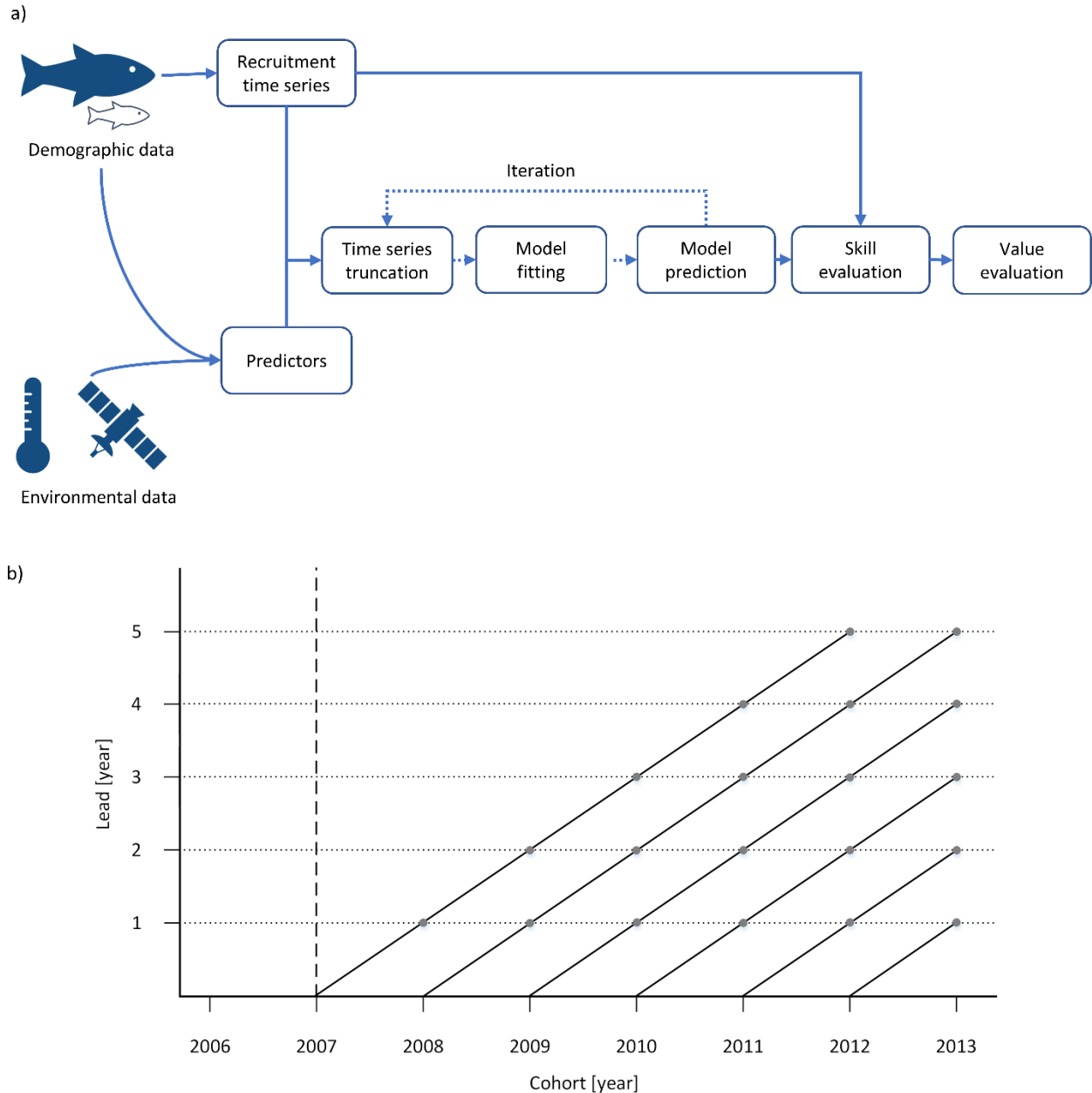
118 differs from other cross-validation approaches (of which it is a subset (Roberts *et al.*, 2017)) and is  
119 important as it directly mimics the way in which recruitment forecasts would be used in an operational  
120 setting. Furthermore more, temporal blocks also remove the potential for the leakage of information  
121 between randomly-selected cross-validation folds, a particularly important issue where there is temporal  
122 structure and autocorrelation in the the time series (as is common in recruitment data). This retrospective  
123 forecasting approach therefore gives a much more realistic assessment of the skill of forecast, and has  
124 been shown to consistently outperform other approaches when forecasting is the goal (Roberts *et al.*,  
125 2017) .

126 The user of the temporal-block approach, however, has two key caveats associated with it. Firstly, the  
127 choice of the initial forecast issue date separating the training and verification blocks represents a tradeoff  
128 between the desire to have as many verifications as possible (and thus the most reliable skill evaluation)  
129 and the need to have sufficient data to train the model on in the first place. This tradeoff is more restrictive  
130 than random cross-validation and will be particularly acute in instances where the length of the time-  
131 series is short: in some cases, there may not be sufficient data to make a reliable skill assessment in this  
132 manner. The exact choice will depend on the characteristics of the system at hand. Secondly, and even  
133 more importantly, care must be taken to avoid inadvertently introducing circular reasoning through the  
134 use of predictors identified by explanatory analyses over the whole time series: such variables will show  
135 skill over the length of the time-series for which they were indentified, but this may not extend into the  
136 future. Ideally, predictors should be based on either generic reasoning (e.g. stock-recruitment  
137 relationships, the match-mismatch hypothesis) or work published prior to the earliest forecast issue date  
138 considered. Alternatively, automatic variable and/or model selection procedures can be incorporated into  
139 the “fit model” part of the framework to allow the identification of skilful predictors for each forecast  
140 issue date.

141 The generic nature of the framework means it can be applied widely: each individual application can and  
142 should vary depending on the specifics of the system being assessed. The recruitment time series used  
143 can be taken from either stock assessment outputs or from a recruitment-index (e.g. from a larval survey).  
144 The selection of predictors is flexible but should be informed by the best available biological knowledge  
145 about the stock (Dickey-Collas *et al.*, 2014b; Subbey *et al.*, 2014) (previous caveats notwithstanding):  
146 stock-specific biomass or demographic indicators, environmental data or other biological parameters (e.g.  
147 prey and predator concentrations) can be incorporated equally. Any modelling approach that produces  
148 predictions can be considered, including classical recruitment models (e.g. Ricker (Ricker, 1954) and  
149 Beverton-Holt (Beverton and Holt, 1957)), statistical and data mining approaches (e.g. generalized  
150 additive models (GAMs) (Hastie and Tibshirani, 1986), empirical dynamic modelling (EDM) (Sugihara *et*  
151 *al.*, 2012) and classifier models (Fernandes *et al.*, 2015)): ensembles of models can also be considered e.g.  
152 combined via multi-model inference (Burnham and Anderson, 2004). Predictions can (and should) be  
153 considered in terms of continuous outputs, probability distributions and/or as categories (i.e.. using a  
154 division into terciles (high, medium, low) based on historical observations). The choice of skill metrics will  
155 be influenced by the nature of the forecast (Jolliffe and Stephenson, 2012) but should include multiple  
156 metrics (Stow *et al.*, 2009; Brun *et al.*, 2016). Skill metrics then form the basis for a quantitative value  
157 assessment, evaluating the expected economic value of following a given forecast. Furthermore, the  
158 framework allows for forecasts of both single stocks or of aggregations of multiple stocks into a single  
159 portfolio forecast, as may be relevant for decision-making across wider-scales (e.g. factories processing  
160 many different species)

161 We illustrate the use of this framework through a worked example focusing on recruitment forecasts of  
162 the lesser sandeel (*Ammodytes marinus*) in the North Sea below.

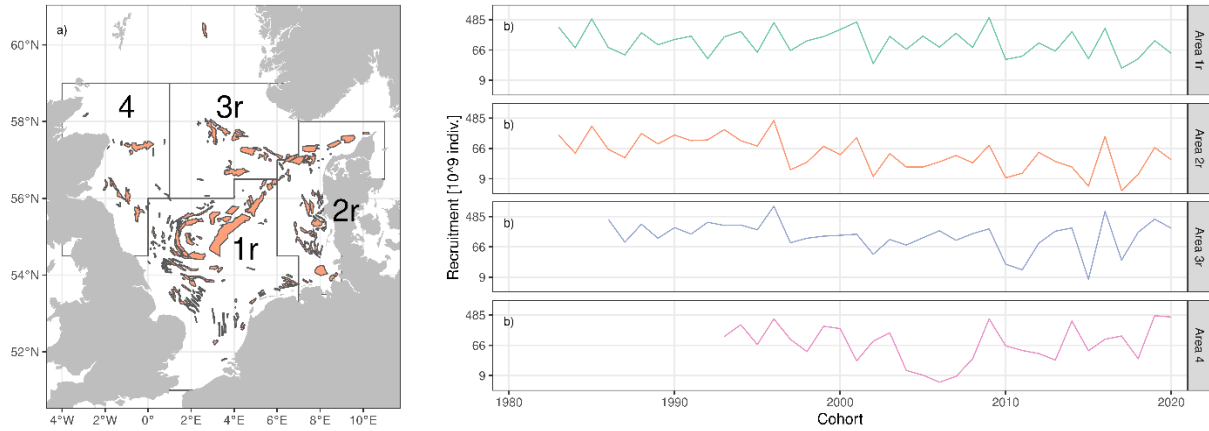




163 **Figure 1 Skill and value assessment framework** a) Overview of the process in the forecasting framework.  
 164 Here, data are extracted and combined into the appropriate modelling data. Afterwards, an iterative  
 165 process of data truncation and model fitting are the basis of all model objects and predictions. These  
 166 predictions contains both the current and the retrospective predictions, which can be used for skill and  
 167 value evaluation b) Schematic of retrospective forecast system used to generate a retrospective forecast  
 168 time series. One time series is generated at each lead time. Dashed line indicates first data cut-off and  
 169 the start of the retrospective forecasting period. Dotted lines indicates the forecast time series at a given  
 170 lead. After the first cut-off each subsequent retrospective forecast will include the previous year's  
 171 observations increasing the size of the model training data set. For each generated retrospective forecast  
 172 time series skill, value and accuracy will be evaluated. Depending on species, stocks, data availability and  
 173 period of interest, the evaluated cohort period and the start of the retrospective analysis can vary.

174 **Sandeel Case Study**

175 The lesser sandeel is a pelagic species of the Ammodytidae family and is one of the most common  
176 sandeels found in the North Sea. Adult lesser sandeel habitats are found in most of the North Sea,  
177 generally distributed across shallow sandy banks (van Deurs et al., 2009, Figure 1a). Particle-tracking  
178 studies and the sedentary state of post-recruitment sandeel (Christensen et al., 2008; Pedersen et al.,  
179 2019) resulted in a division into 7 different individually managed North Sea sandeel stocks. Analytical  
180 stock assessments are done in management areas 1r, 2r, 3r and 4 (see Figure 2a), while the remaining  
181 three stocks are considered data poor. Sandeel is seen as one of the main links between primary  
182 production and the higher trophic levels in the North Sea for both larger piscivorous fish (e.g. cod and  
183 haddock) and seabirds (Eliassen *et al.*, 2011). The lesser sandeel has historically supported a large fishery,  
184 which has seen a large decline in recent years (Dickey-Collas *et al.*, 2014b). Due to the importance of the  
185 species, recruitment to these stocks is well studied (Arnott and Ruxton, 2002; van Deurs *et al.*, 2009;  
186 Eigaard *et al.*, 2014; Lindegren *et al.*, 2018). In the south western part of the North Sea (i.e. management  
187 area 1r), sandeel shows signs of being influenced negatively by temperature, while the abundance of the  
188 main prey, *Calanus finmarchus*, has a positive influence (Arnott and Ruxton, 2002; Lindegren *et al.*,  
189 2018). Density dependence has also been found to be an important driver, where competition with  
190 young adults and juveniles has a negative effect on recruitment (van Deurs *et al.*, 2009). Currently, stock  
191 assessment uses a geometric mean for recruitment predictions (ICES, 2018). These geometric means will  
192 be used as continuous reference models during skill evaluation.



193 **Figure 2 Study area and data** a) Map of the North Sea showing the four management areas of sandeel  
194 assessed analytically. Sandy habitat banks, the predominant sandeel habitat, are shown in orange. b)  
195 Recruitment time series for the four sandeel stocks from the official ICES stock assessment. Dashed-  
196 horizontal lines mark the delineation of the upper and lower terciles for each stock.

197

198 **Data**

199 Operational forecasts require data to be available at the time of the forecast, potentially excluding some  
200 potentially relevant predictors. For example, estimates of zooplankton prey, *Calanus finmarchicus* and  
201 *Temora longicornis* have been used in other explanatory studies (Arnott and Ruxton, 2002; van Deurs *et*  
202 *al.*, 2009; Lindegren *et al.*, 2018) but are only available with 1-2 years delay, and are therefore of limited  
203 value in forecasting recruitment in this stock in an operational setting. We focus our analyses on data that  
204 are available with a maximum of a few months delay. An overview of the data employed is provided in  
205 Table 1 and the complete time series are found in Figure S1.

206 **Assessment data**

207 Assessment data used for sandeel modelling is obtained from official ICES advice, based on the stochastic  
208 multi-species assessment model, SMS (Pedersen *et al.*, 1999). The SMS model is run in a single-stock mode  
209 for sandeel assessments, and integrates data on catches, catch effort, maturity, weight, fishing mortality  
210 and natural mortality at a given age (ICES, 2018). All stock assessment data are the current (2021)  
211 assessments provided by ICES for area 1r, 2r, 3r and 4 (Figure 2b), where recruits are treated at age 0.  
212 From the assessment data, 4 demographic variables are extracted, consisting of spawning stock biomass  
213 (SSB), total stock biomass (TSB), number of individuals (SumN) and number of one-year olds (N1). This  
214 allows for different types of interactions between the demography, including density dependence and  
215 SSB impact on recruitment. All demographic data are log-transformed before use in modelling and  
216 converted to log-anomalies (relative to the average log-value over the full time series for each stock).

217 **Environmental data**

218 High resolution spatial sea surface temperature data is gathered from the Optimum Interpolation Sea  
219 Surface Temperature (OISST) product (Banzon *et al.*, 2016). The product is a 0.25° x 0.25° global daily sea  
220 surface temperature (SST) data set on a regular grid. Noting that adult sandeel are bound to specific banks

221 (Christensen *et al.*, 2008), we produced daily average temperatures over the banks in each stock area, and  
222 then averaged temporally over quarters as follows: P3 and P4 represents the temperature anomalies  
223 experienced by the adult sandeel from July to December before and during spawning (i.e. the  
224 temperatures experienced by the spawners just before spawning). Q1, Q2, Q3 and Q4 are SST anomalies  
225 experienced during the egg, larval and juvenile stages from January to December for a given cohort. All  
226 extracted temperatures were converted to anomalies from the average (climatology) over the complete  
227 SST time series period (1983 to 2020) prior to use in modelling.

## 228 **Models**

229 Here we use generalised additive models (GAM) as the basis for generating predictions, with model  
230 variable selection based on a multi-model inference approach. An advantage of the GAM approach is it's  
231 semi-parametric nature that allows for arbitrary but smooth responses. We exploited this feature to  
232 incorporated a cohort-based time-varying smoother to allow for changes in the underlying productivity  
233 (e.g. due to unquantified variables). This approach allows non-stationarity and systematic shifts in  
234 recruitment patterns that would otherwise not be accounted for. For in-depth model descriptions, see  
235 supplementary methods.

236 The total of 11 candidate variables (Table 1) give a total of 2048 possible combinations that could be  
237 considered. However, in order to minimize risk of overfitting due to both collinearity between model  
238 parameters and the short time-series, predictors are split into three groups (as shown in *Table 1*) based  
239 on an exploratory analysis of collinearity (i.e. environmental, demographic and other predictors).

240 Models in the ensemble that incorporated more than one variable in a given group were excluded,  
241 giving a total of 819 candidate model structures to be considered.

242 Following the retrospective-forecasting and time-blocking approach proposed in this framework, (Figure  
243 1), models were first trained on all data up to a cut-off point and the small-sample Aikaike Information

244 Criteria (AICc) calculated and converted to model weights (Anderson, 2008). Each model was then used  
245 to predict the distribution of expected recruitment values for each cohort in the second, verification  
246 block. The individual model posterior predictions were then combined into an ensemble predictive  
247 distribution, with the contribution of each model to the ensemble prediction being determined by the  
248 AICc weights. Probabilistic categories (i.e. high, medium and low recruitment) and the expected value  
249 (mean across the distribution) were then generated from this ensemble predictive distribution. This  
250 process was repeated by moving the cut-off point (forecast issue date) forwards by one year, creating a  
251 forecast various lead times (Figure 1b).

252 We evaluated forecast issue dates from 2007-2020, giving a total of 14 forecasts to evaluate: earlier  
253 first-forecast dates struck problems with model stability due to the short time series in area 4 (starting  
254 from 1993). We focused on the first forecast (one cohort ahead) here, as this is the most relevant to  
255 both the management of the stock and to the associated fishing industry.

256

257 *Table 1 List of all variables considered and the rationale behind. The parameterisation of each variable in*  
 258 *the model is also shown, with s() indicates the use of a spline-smoother and other terms indicating the*  
 259 *incorporation of that term as a linear response term.*

<b>Variable</b>	<b>Description</b>	<b>Rationale</b>	<b>Parameterisation</b>
<b>Demographic explanatory variables</b>			
<i>SSB</i>	Spawning stock biomass	Adult biomass that determines amount of eggs spawned	s(log(SSB))
<i>N1</i>	Number of 1-year olds	Number of individuals at age 1 inducing a density dependence	log(N1)
<i>SumN</i>	Number of individuals	Entire sandeel population, inducing density dependence	log(SumN)
<i>TSB</i>	Total stock biomass	Combination of all of the above	log(TSB)
<b>Enviromental explanatory variables</b>			
<i>P3</i>	Jul-Sep temperatures	Temperatures experienced by the adults prior to spawning	P3
<i>P4</i>	Oct-Dec temperatures	Temperatures experienced by the adults prior to / during spawning	P4
<i>Q1</i>	Jan-Mar temperatures	Temperature experienced during egg development	Q1
<i>Q2</i>	Apr-Jun temperature	Temperature experienced by larvae during pelagic drift phase	Q2
<i>Q3</i>	Jul-Sep temperature	Temperature experienced by post-settlement juveniles	Q3
<i>Q4</i>	Oct-Dec temperature	Temperature experienced by post-settlement juveniles	Q4
<b>Other explanatory variables</b>			
<i>Cohort</i>	Cohort year	Included to allow time-variation in the mean productivity of the stock due to systematic shifts in other unquantified variables	s(Cohort)

261 **Skill metrics**

262 Multiple performance metrics are used to assess the retrospective forecasts (*Table 2*), including both  
263 continuous and categorical skill evaluations (Stow *et al.*, 2009; Jolliffe and Stephenson, 2012; Brun *et al.*,  
264 2016). Continuous skill uses the mean prediction for a root-mean-square error (RMSE) analysis, giving  
265 indications of the accuracy of the forecast. Continuous forecasts can use the mean-squared-error skill  
266 score (MSESS) to directly compare the forecast with a reference forecast. The categorical forecasts (high,  
267 medium and low) are analysed using the hit rate (H), false alarm rate (F) and true skill score (TSS). Using a  
268 combination will quantify both the accuracy of the forecast and forecast performance in each tercile  
269 (Murphy, 1969).

270 Reference forecasts were selected according to current stock assessment practices: in this way, it was  
271 immediately apparent if the forecast outperforms existing procedures. For the sandeel, the official ICES  
272 sandeel advice uses either the 10-years moving geometric mean (Area 2r and 4) or the geometric mean  
273 of the full time series (Area 1r and 3r) (ICES, 2018). These models are selected as reference forecasts in  
274 the MSESS. The skill score ranges from negative infinity to 1, effectively comparing the performance gains  
275 from using a given forecast compared to the reference. For categorical forecasts, the reference forecast  
276 is selected to be random guessing (33% correct) baseline, both for True Skill Score (TSS) and Ranked  
277 Probability Skill Score (RPSS).

278



279 *Table 2 Performance metrics used to evaluate the skill of the forecast system. These values are calculated*  
 280 *over all retrospective forecasts at a given lead time. Continuous skill evaluation is also performed for*  
 281 *reference forecasts, while the categorical and binary forecast evaluation are only calculated for*  
 282 *probabilistic forecasts (Murphy, 1969). MSE contains the mean of the difference between the forecasted*  
 283 *(F) and observed (O). The hit rate (H) consists of the proportion of correct forecasts (i.e. true positives (TP)*  
 284 *and true negatives (TN)), while false alarm rate (F) is the proportion of incorrect forecasts (i.e. false*  
 285 *positives (FP) and false negatives (FN)).*

Name of Forecast Quality Measure	Definition	Range	Application
Mean square error (MSE)	$\frac{1}{n} \sum_{i=1}^n (F_i - O_i)$	[0,inf]	Cont.
Mean square error skill score (MSESS)	$1 - \frac{MSE}{MSE_{reference}}$	[-inf,1]	Cont.
Root-mean-square error (RMSE)	$\sqrt{MSE}$	[0,inf]	Cont.
Proportion correct / Hit rate	$H = \frac{TP + TN}{(TP + TN) + (FP + FN)}$	[0,1]	Cat.
False alarm rate	$F = \frac{FP + FN}{(TP + TN) + (FP + FN)}$	[0,1]	Cat.
True Skill Score (TSS)	$TSS = H - F$	[-1,1]	Cat.

286

287 *Table 3 a) Confusion matrix generated from the retrospective forecasts at a given lead. Constructed from*  
 288 *the sum of observed positives (event occurred) and observed negatives (event didn't occur) with*  
 289 *corresponding predicted positives (predict event occurred) and predictive negatives (predicted event not*  
 290 *to occur). This results in a matrix of true positive (TP), false positives (FP), false negatives (FN) and true*  
 291 *negatives (TN). For recruitment predictions, a TP is when the forecasting system correctly predicts the*  
 292 *observed tercile, while a FP is when the system predicts a given tercile, which is not observed. For negative*  
 293 *events this is reversed, i.e. FN the tercile is observed while the forecast system doesn't predict it and TN is*  
 294 *when the tercile is not observed and the system doesn't predict it. b) Cost matrix used to calculate the*  
 295 *value of the forecast system. Here a cost (C) is associated with a precaution and a loss is associated with*  
 296 *not taking the precaution and the event occurring.*

a) Contingency matrix

	Observed P	Observed N
Predict P	TP	FP
Predict N	FN	TN

b) Cost matrix

	Event occurs	Event does not occur
Precaution taken	C	C
Precaution not taken	L	0

297

## 298 Forecast value

299 We assess the value of the forecasts using a Richardson cost-loss decision model (Richardson, 2000).  
300 Simple economic models, as used here, are widely used in the climate services sector (Pope *et al.*, 2019)  
301 to quantify value of e.g. seasonal forecast systems, and provide an intuitive metric for users (Murphy,  
302 1976b). Briefly, the model considers the economic impacts of a particular event that is being forecast  
303 (e.g. poor recruitment), and the loss (L) that the user could potentially incur. However, the user also has  
304 the ability to avert these losses by implementing precautionary mitigation actions (e.g. based on a  
305 forecast), but doing so also incurs a cost (C) (e.g. mothballing processing plants). These two dimensions  
306 (i.e. whether the event occurs, and whether the user takes a precaution) each have two outcomes, and  
307 therefore form a 2x2 cost matrix (Jolliffe & Stephenson, 2012, see Table 3b). Combining this set of costs  
308 with the properties of the forecast system characterised by the contingency matrix (Table 3a) allows the  
309 expected expense over the long-term (E) to be calculated when the forecast is always ( $E_{forecast}$ ) or never  
310 ( $E_{reference}$ ) followed. The value (V) of the realised forecast system can then be calculated relative to a  
311 perfect forecast system as :

$$312 \quad V = \frac{E_{reference} - E_{forecast}}{E_{reference} - E_{perfect}} \quad (1)$$

313 The value of the forecast system, V, is expressed as a non-dimensional number less than 1 and varies as a  
314 function of the cost-loss ratio (C/L) of a given user (Richardson, 2000, see eq. 1).

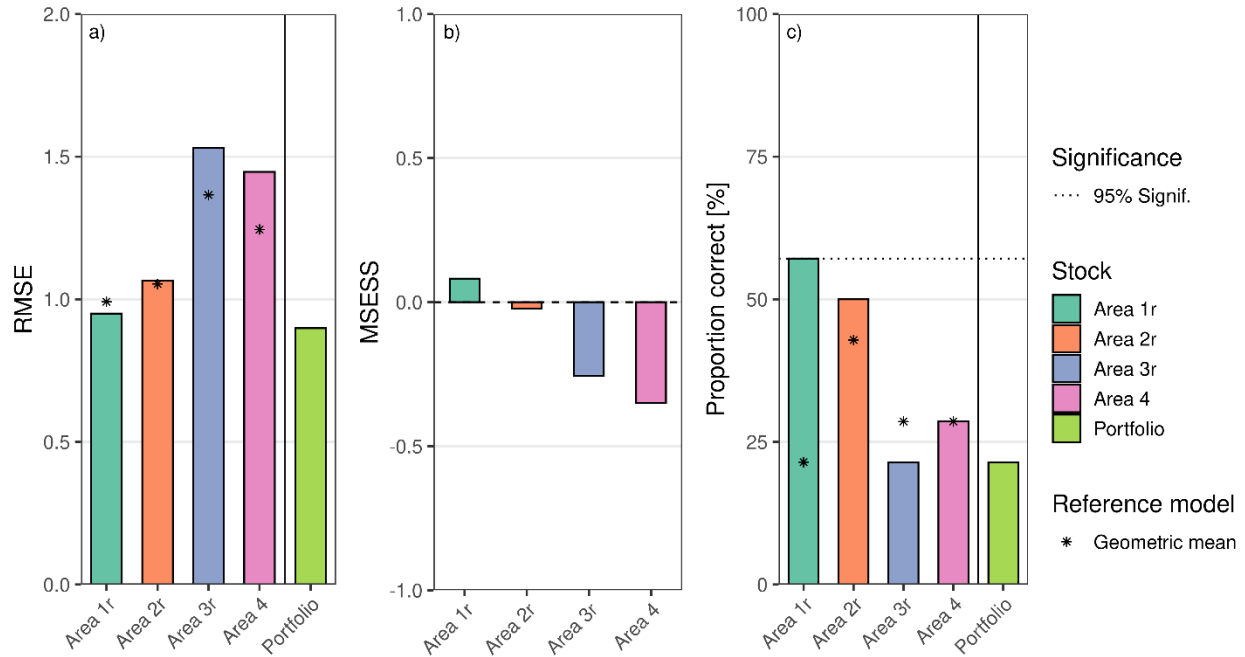
315 We necessarily extend this analysis to account for the (relatively) small sample size associated with our  
316 set of retrospective forecasts and therefore estimate the uncertainties in the value. We model the  
317 retrospective contingency table (Table 3a) using a Bayesian multinomial model implemented in Stan  
318 (Stan Development Team, 2020) to estimate the vector of true probabilities  $\mathbf{p} = \{p_{TP}, p_{FP}, p_{FN}, p_{TN}\}$  of  
319 each quadrant of the contingency table. The posterior predictive distribution of  $\mathbf{p}$  was then sampled

320 4000 times and used to construct a corresponding large set of contingency tables and therefore the  
321 statistical distribution of the forecast system value,  $V$ .

## 322 **Results**

323 Assessment of the predictions is presented at a forecast lead of one cohort beyond the final year of the  
324 assessment, mimicking potential operational usage in these stocks. We find that the stocks in area 1r and  
325 2r have the highest continuous forecast accuracy, while areas 3r and 4 show higher RMSEs (Figure 3a):  
326 this dichotomy closely parallels the lengths of the time series of each area (areas 3r and 4 being  
327 appreciably shorter) and we hypothesize that the reduced amount of training data may limit the forecast  
328 skill. Furthermore, the assessment of area 1r is widely perceived as being the most reliable of the four:  
329 the poor performance in areas 3 and 4 in particular may be due to the poor quality of the assessment as  
330 much as the poor quality of the forecast. The portfolio forecast, on the other hand, has the highest overall  
331 accuracy, showing that the aggregation of predictions can lower the RMSE, highlighting the smoothing  
332 effect associated with aggregating noisy data sets.

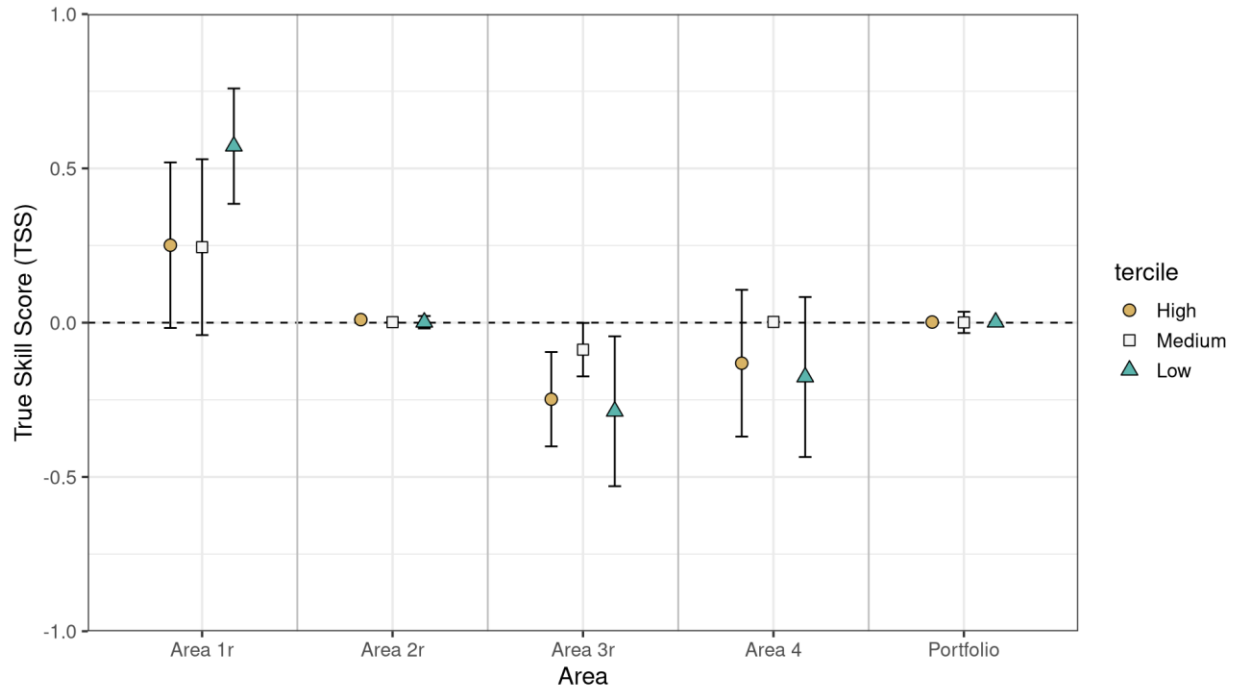
333 Comparing our forecasts against the existing models used in the assessment of this stock (geometric  
334 mean) places their skill in context. In management area 1r, the continuous forecast accuracy is better than  
335 these reference models (Figure 3a), giving a positive mean-squared error skill score (MSESS) (Figure 3b).  
336 The performance of Area 2r is on a par with the reference model, while area 3r and 4 both show a negative  
337 MSESS at lead 1, indicating that the forecast model ensemble would not be an improvement over the  
338 geometric mean reference model when used as a continuous forecast.



339

340 **Figure 3 Recruitment forecasts outperform reference forecasts in some cases** a) Root-mean-squared  
 341 error of the different management areas and the portfolio forecast at lead 1. Area 1r and 2r show highest  
 342 accuracy of individual forecasts, while the portfolio is the overall most accurate, indicating the presence  
 343 of the portfolio effect. Stars show the reference geometric mean RMSE for the individual management  
 344 areas. b) Mean-squared error skill score of the individual forecast products for lead 1. Official recruitment  
 345 prediction model is used as a reference model. Here area 1r and 2r shows better or equal performance to  
 346 the reference models, while area 3 and 4 has a negative skill score. c) Hit rate of the different management  
 347 areas indicating the percentage of correct retrospective forecasts at lead 1. Dashed line indicates the 95<sup>th</sup>  
 348 percentile level of the random guessing reference forecast. Area 1r are significantly better than random  
 349 guessing at 57% hit rate, while area 2r are borderline significant with 50% hitrate. Area 3r, 4 and the  
 350 portfolio shows large drop-offs in hit rate with hit rates below 30%. Stars show the reference geometric  
 351 mean hit rate.

352



353 **Figure 4 Categorical recruitment forecasts show skill in some areas.** Model skill at lead 1 is represented  
354 as the True (Peirce) Skill Score (TSS), which ranges between +1 and -1, and has a value of 1 for perfect skill,  
355 and 0 for random guessing (black dashed line). Negative values indicates perverse forecast. The 95%  
356 confidence interval for the estimated skill score are shown as error bars on each of the points. Recruitment  
357 stocks are shown, with shapes indicating the corresponding recruitment tercile. A positive TSS is seen for  
358 all recruitment terciles in area 1r, while all other models shows utility close to or worse than random  
359 guessing.

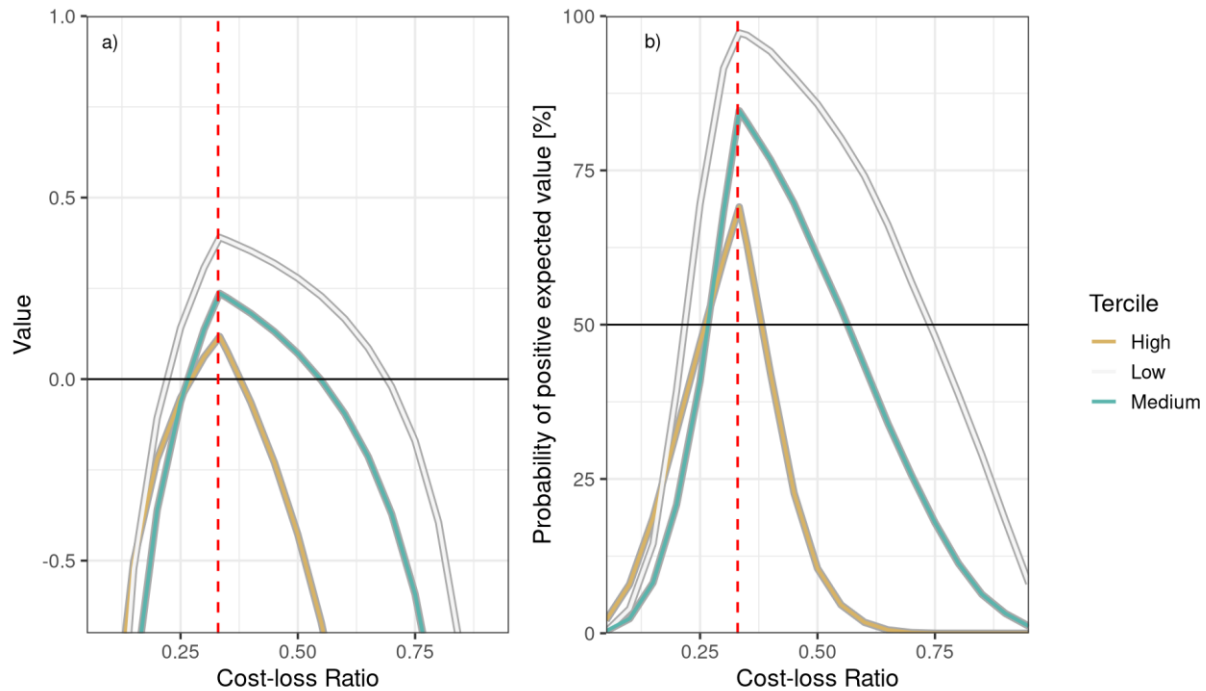
360

361 The categorical performance of the forecast models is also broadly similar. Hit rate metrics (how often the  
362 system correctly forecasts high, medium or low recruitment) also shows best results in management area  
363 1r, with 57% correct (Figure 3c), outperforming the outperforming expected 33% correct associated with  
364 the random guessing of terciles ( $p=0.02$ , one-tailed test). Area 2r sees a hit rate of 50% correct ( $p=0.06$ ,  
365 one-tailed test), significant at the 90% level. A large drop off in hit rate is seen in area 3r and 4 (respectively  
366 at 21% and 28%), where performance is not significantly better than random guessing ( $p=0.74$ , and  
367  $p=0.52$ , one-tailed tests). The portfolio categorical forecast, on the other hand, performs poorly and is not  
368 significantly better than random guessing at a 21% proportion correct ( $p=0.74$ , one-tailed test): while  
369 aggregating improves the performance of continuous forecasts, it clearly deteriorates categorical  
370 forecasts.

371 Further insight into the forecast system can be gained by examining the skill of predicting individual  
372 terciles. The true-skill score (TSS) metric combines the specificity (true-positive rate) with the sensitivity  
373 (true-negative rate) for a categorical forecast and is applied here to each tercile in turn. The TSS indicates  
374 area 1r being the only management area where the model can reliably differentiate all three categories  
375 (Figure 4), consistently outperforming random guessing (i.e. where  $TSS=0$ ). Most other areas do not show  
376 a significant ability to differentiate, either due to the small sample size or poor model skill. For example,  
377 area 2r's TSS is not significantly different from zero for all terciles, in part due to the relatively low  
378 recruitment seen in the stock in recent years, affecting the ability of the TSS metric to quantify the forecast  
379 skill. Areas 3r and 4 have negative or zero skill scores in all categories, likely due to the aforementioned  
380 poor quality of these assessments propagating into these forecasts and resulting in a wide prediction  
381 distribution. The portfolio forecast shows similar TSS values to area 2r, with no categories reaching levels  
382 where the system is able to correctly distinguishing between terciles.

383 We assessed the value of the forecast for all areas. The cost-loss decision model for Area 1r (Figure 5)  
384 shows positive values in all forecast categories, with especially the low recruitment prediction showing

385 the highest value over a broad range of cost/loss ratios (Figure 5a). All categories peak at a cost-loss  
386 ratio of 0.33, as is expected from theoretical analyses of this model (Jolliffe and Stephenson, 2012). We  
387 account for the small sample size and propagate the uncertainty that it creates into the forecast value  
388 by estimating the probability of a positive expected value for a given cost/loss ratio (Figure 5b): this  
389 metric provides decision makers with an indicator when using the forecast will lead to a positive  
390 economic return. Here the peak is still seen at a cost-loss ratio of 0.33, where all categories have above  
391 65% probability of a positive expected long-term value. Following the low recruitment forecast for this  
392 cost-loss ratio (i.e. 0.33) will result in a 96% probability of positive value from the forecast, but  
393 probabilities above 50% are also seen across a wide range of cost-loss ratios. While area 2r, 3r and 4  
394 generally can't provide the same levels of value, area 2r could prove valuable when following the high  
395 forecast (Figure S3).



396 **Figure 5 Positive economic value is seen in area 1r recruitment forecasts.** Long-term value of a cost-loss  
397 decision model in area 1r, simulated from a multinomial confusion matrix model. a) Tercile divided value  
398 given cost-loss ratios. Solid line indicates zero value. Positive value is seen in all terciles, peaking at a cost-  
399 loss ratio of 0.33. Most value can be gained by following the low tercile forecast, which corresponds with  
400 the highest observed TSS. b) Tercile divided probability of a positive expected value. Calculated from a  
401 Bayesian posterior distribution, indicating the probability of drawing a positive value at a given cost-loss  
402 ratio. Peak probability is seen at cost-loss ratio of 0.33, where all terciles shows above 65% probability of  
403 a positive expected value.

404



## 405 **Discussion**

406 Here we present a framework for robustly assessing the skill and value of recruitment predictions in a way  
407 that is relevant to their use in an operational setting. The case study that we have examined, for four  
408 sandeel stocks in the North Sea, illustrates several important conceptual points that deserve particular  
409 attention.

410 Firstly, we show the importance of assessing a forecast system with multiple metrics. While in-sample  
411 performance and explanatory metrics are good for finding correlations (and thereby highlighting possible  
412 causality), the assessment of predictive skill is quite different and should primarily be shaped by the needs  
413 of the forecast user. For example, we identify an overall high forecast accuracy in area 2 (RMSE in Figure  
414 3), but the ability to distinguish between the two lower terciles is poor (Figure 4). Stock assessors may  
415 focus on the MSESS as a criteria for uptake, while industry might be more interested in performance in a  
416 specific category (e.g. ability to forecast poor year classes) or long-term economic value. Furthermore, the  
417 value of a forecast to users within the same sector (e.g. two different fish processing plants) may differ  
418 due to differences in their underlying risk profile (i.e. cost-loss ratio) such that while the forecast system  
419 may be advantageous for one user, it may not be of use to another. Understanding the decision-making  
420 needs of the user is therefore essential to the production of a good forecast (Murphy, 1993; Payne *et al.*,  
421 2017).

422 While the application of the cost-loss model to estimate forecast value has a clear interpretation in a  
423 commercial context, it is less clear how relevant this approach is to fisheries management. Here, cost-loss  
424 decision models encapsulate both the costs and losses associated with correctly and incorrectly  
425 forecasting recruitment. These ideas can be relevant to fisheries management, as the managers can use  
426 this knowledge as the basis of the forecast evaluation, assigning value on e.g. true positives versus false  
427 positives. This allows managers and users to understand how the forecast can be incorporated, and how

428 forecasts can and should be used in the management of a given stock. While not an economic gain, the  
429 value metric of a forecast can be used to assess and manage stocks sustainably, providing the managers  
430 with the tools to properly assess how to incorporate forecasts into decision making. Our demonstration of  
431 the framework here is based on the use of recruitment estimates directly from the stock assessment, as  
432 is still common in the field. It is nevertheless important to remember that these data are estimates that  
433 are also uncertain (Brooks and Deroba, 2015). The framework presented here has the ability, however,  
434 incorporate a more robust treatment of such uncertainties. For example, uncertainty estimates (e.g. in  
435 recruitment) can be incorporated directly into forecast model if desired. Retrospective biases in the stock  
436 assessment incorporated into the model fitting procedure by e.g. fitting the forecast model to stock-  
437 assessment outputs based on a model up to 2007, and then predicting forward in time from there. While  
438 such an approach would be ideologically cleaner, it was not possible here due to technical challenges in  
439 producing a sufficient number of retrospective assessments for these stocks. A further extension would  
440 be to incorporate the recruitment forecast model directly into the stock assessment model, thereby  
441 making a seamless assessment and recruitment prediction system. Regardless of the approach, the  
442 framework presented can adapt to both the technical limitations of the system being studied, and  
443 changing norms in the approach to this issue.

444 Finally, our results for North Sea sandeel show that our understanding of recruitment predictability needs  
445 to be re-assessed. Contrary to the wide-spread belief that recruitment can't be forecast, we have shown  
446 in a setting that directly mirrors operational useage that skilful and valuable recruitment forecasts can be  
447 made. Shifting the way that we assessment recruitment skill from an explanatory to predictive setting  
448 greatly increases the confidence in, and transparency of, these results, and paves the way for their direct  
449 up-take in decision making. Furthermore, taking the next step of assessing the value of these forecasts  
450 gives a more nuanced view that is directly relevant to decision-makers, particularly in the commercial

451 sector. These results therefore open the way for a new paradigm in addressing this long-running, but  
452 fundamental question in fisheries management.

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### 458 **Data availability**

459 The data that support the findings of this study are available from the corresponding author upon  
460 reasonable request.

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